| Id | 1 |
|----------|--|
| Question | It is given that, $P(A B)=0.4$ and $P(A B^c)=0.6$ then |
| A | P(A) = 0.5 |
| В | $0 \le P(A) \le 1$ |
| С | $0.4 \le P(A) \le 0.6$ |
| D | $0.6 \le P(A) \le 1$ |
| Answer | C |

| Id | 2 |
|----------|--|
| Question | E and F are independent events such that, the probability of at least one occur is 1/3 and |
| | the probability that E occurs but F does not occurs in 1/9. Then P(F) is, |
| A | 4/9 |
| В | 2/9 |
| С | 1/9 |
| D | 1/3 |
| Answer | В |

| Id | 3 |
|----------|--|
| Question | Suppose that, the distribution of r.v. X is given by, |
| | P(X=-1)=1/6=P(X=4) and $P(X=0)=1/3=P(X=2)$ then, find $E(X/X+2)$ |
| A | 1/6 |
| В | 1/9 |
| С | 1/18 |
| D | 1/36 |
| Answer | В |

| Id | 4 |
|----------|--|
| Question | Let ~ N(1,4). Given that, $\frac{X^2}{4} - \frac{X}{2} + \frac{K}{4}$ is distributed as Chi-square with one degree |
| | of freedom. The value of K is |
| A | 0 |
| В | 4 |
| С | 1.414 |
| D | 1 |
| Answer | D |

| Id | 5 |
|----------|--|
| Question | Let X have the distribution function $ \begin{array}{c c} 0 & \text{if } X < -1 \end{array} $ |
| | $F(X) = \begin{cases} 0, & \text{if } X < -1\\ \frac{2+X}{4}, & \text{if } -1 \le X \le 1\\ 1, & \text{if } X \ge 1 \end{cases}$ |
| | The $V(X)$ is, |
| A | 1/2 |
| В | 1/4 |
| С | 1 |
| D | 2/3 |
| Answer | A |

| Id | 6 |
|----------|---|
| Question | If x has a pdf, $f(x)=2(1-x)$, $0 \le x \le 1$. What is the median of the distribution. |
| A | $1+\frac{1}{\sqrt{2}}$ |
| В | $1-\frac{1}{\sqrt{2}}$ |
| С | $\frac{1}{\sqrt{2}}$ |
| D | $\sqrt{2}-1$ |
| Answer | В |

| Id | 7 |
|----------|--|
| Question | The characteristic function of $U(-\alpha, \alpha)$ is |
| A | $\frac{1}{\alpha t}\cos(t\alpha)$ |
| В | $\alpha t \sin(t \alpha)$ |
| С | $\frac{1}{\alpha t}\sin(t\alpha)$ |
| D | $\alpha t \cos(t \alpha)$ |
| Answer | C |

| Id | 8 |
|----------|---|
| Question | If $b_{yx} = 0.4$ and $b_{xy} = 0.16$ are the regression coefficient on Y on X and X on Y |
| | respectively, then the correlation coefficient between X and Y, |
| A | 0.08 |
| В | -0.088 |
| С | 0.0064 |
| D | None of these |
| Answer | D |

| Id | 9 |
|----------|---|
| Question | If Markov chain is finite then, |
| A | There exist at least one state which is transient |
| В | There exist at least one state which is recurrent |
| С | All states are absorbing or reflecting barriers |
| D | All states are either recurrent or transient |
| Answer | В |

| Id | 10 |
|----------|---|
| Question | If random sample of size n is chosen from a population with pdf, |
| | $f(x,\theta) = \begin{cases} \frac{1}{2}e^{-(x-\theta)}; x \ge \theta \\ \frac{1}{2}e^{(x-\theta)}; x < \theta \end{cases}$ Then the m.l.e. of θ is, |
| A | Mean of the sample |
| В | Standard deviation of the sample |
| С | Median |
| D | $X_{(n)}$ |
| Answer | C |

| Id | 11 |
|----------|--|
| Question | For testing $H_0: \mu=0$ against $H_1: \mu>0$ on the basis of a random sample |
| | X_1, \dots, X_n a proposed tests rejects H_0 if and only if the number of positive observations is too large. The test is: |
| A | UMP for the problem |
| В | Not UMP but UMPU for the problem |
| С | Not UMP but unbiased for the problem |
| D | Not UMP and not unbiased for the problem |
| Answer | D |

| Id | 12 |
|----------|---|
| Question | If T_1 and T_2 are unbiased estimators of θ and θ^2 $(0 < \theta < 1)$ and T is |
| | sufficient statistics, then $E(T_1/T) - E(T_2/T)$ is: |
| A | The minimum variance unbiased estimator of θ |
| В | Always the minimum variance unbiased estimator of $\theta(1-\theta)$ |
| С | Always an unbiased estimator of $\theta(1-\theta)$, which has variance not exceeding that of |
| | $T_1 - T_2$ |
| D | Not and unbiased estimator of $\theta(1-\theta)$ |
| Answer | С |

| Id | 13 |
|----------|--|
| Question | If X_1, \dots, X_n is a random sample from Poisson distribution with mean λ , The |
| | Cramer – Rao lower bound to the variance of any unbiased estimator of λ is given by: |
| A | λ^2/n |
| В | λ/η |
| С | $\lambda^{1/2}/n$ |
| D | $e^{-\lambda}/n$ |
| Answer | В |

| Id | 14 |
|----------|---|
| Question | Let X_1, X_2, \dots, X_n be a sequence of i.i.d. r.v. With $P(X_i = -1) = P(X_i = 1) = 1/2$. |
| | Suppose the standard normal variable Z, $P[-1 < Z \le 0.01] = 0.08$. If $S_n = \sum_{i=1}^{n^2} X_i$ then $\lim_{n \to \infty} P(S_n > n) = 2$ |
| | $\lim_{n \to \infty} P(S_n > \frac{n}{10}) = ?$ |
| A | 0.42 |
| В | 0.46 |
| С | 0.50 |
| D | 0.54 |
| Answer | В |

| Id | 15 |
|----------|---|
| Question | For which of the following distributions, the WLLN does not hold? |
| A | Normal |
| В | Gamma |
| С | Beta |
| D | Cauchy |
| Answer | D |

| Id | 16 |
|----------|--|
| Question | Let X be r.v. Such that, $E[X]=E[X^2]=1$. Then $E[X^{100}]$ is: |
| A | 0 |
| В | 1 |
| С | 2 ¹⁰⁰ |
| D | $2^{100}+1$ |
| Answer | В |

| Id | 17 |
|----------|---|
| Question | Suppose X is real valued r.v. Which of the following values cannot be attained by the |
| | following values cannot be attained by $E[X]$ and $E[X^2]$ respectively? |
| A | 0 and 1 |
| В | 2 and 3 |
| С | 1/2 and 1/3 |
| D | 2 and 5 |
| Answer | В |

| Id | 18 |
|----------|--|
| Question | Suppose, $X \sim \text{Exp}(1/\lambda)$, where $\lambda > 0$. For testing the hypothesis $H_0: \lambda = 3$ Vs. |
| | $H_1: \lambda = 5$, a test is given as, "Reject H_0 is $X \ge 4.5$ " The probability of type-I error and power of this test are respectively? |
| A | 0.1353 and 0.4466 |
| | |
| В | 0.1827 and 0.379 |
| C | 0.2021 and 0.4493 |
| D | 0.2231 and 0.4066 |
| Answer | D |

| Id | 19 |
|----------|---|
| Question | Let $\{X_i\}$ be a sequence of independent Poisson (λ) variables and let |
| | $W_n = \frac{1}{n} \sum_{i=1}^{n^2} X_i$ Then limiting distribution of $\sqrt{n}(W_n - \lambda)$ is the normal distribution |
| | with mean zero and variance is given by: |
| A | 1 |
| В | $\sqrt{\lambda}$ |
| С | λ |
| D | λ^2 |
| Answer | C |

| Id | 20 |
|----------|---|
| Question | Which of the following distribution have Coefficient of variation is one? |
| A | Normal |
| В | Poisson |
| С | Gamma |
| D | Exponential |
| Answer | D |

| Id | 21 |
|----------|--|
| Question | Consider a LPP: Maximize $Z=4X_1+3X_2$, subject to, $X_1+X_2\geq 3$, $X_1-X_2\geq 2$, |
| | $X_1, X_2 \ge 0$. An optimal solution of its dual problem is given by, |
| A | $X_1 = 0, X_2 = 3$ |
| В | $X_1 = 1, X_2 = 1$ |
| С | $X_1 = 2, X_2 = 0$ |
| D | Unbounded solution |
| Answer | D |

| Id | 22 |
|----------|---|
| Question | Branch and bound method is |
| A | Not used to solve any kind of programming problem |
| В | Cannot divide the feasible region into smaller parts |
| С | Cannot employ in problems when there are finite numbers of solution |
| D | Standard method and can applied differently for different problems |
| Answer | A |

| Id | 23 |
|----------|--|
| Question | The quadratic form $X^T QX$ is said to be positive semi-definite, if |
| A | $X^T Q X > 0$ |
| В | $X^T Q X < 0$ |
| С | $X^T Q X \ge 0$ |
| D | $X^T Q X \le 0$ |
| Answer | C |

| Id | 24 |
|----------|---|
| Question | In the EOQ problem without shortages with finite replenishment, the optimum lot size is |
| A | $\sqrt{\frac{2DC_S}{C_1(1-\frac{1}{k})}}$ |
| В | $\sqrt{\frac{2DC_s}{C_1(1-\frac{r}{k})}}$ |
| С | $\sqrt{\frac{2DC_S}{C(1-\frac{r}{k})}}$ |
| D | $\sqrt{\frac{DC_S}{IC(1-\frac{r}{k})}}$ |
| Answer | В |

| Id | 25 |
|----------|--|
| Question | In $\{(M M 1):(N FIFO)\}$ queuing model, the average waiting time in the queue is, |
| A | $E(w) = E(v) - 1/\mu$ |
| В | $E(w) = E(v) - 1/\lambda$ |
| С | $E(w) = E(n) - 1/\mu$ |
| D | $E(w) = E(n) + 1/\lambda$ |
| Answer | A |

| Id | 26 |
|----------|---|
| Question | Consider a BIBD with λ , a, r and k as the design parameters in the usual notation. |
| | Which of the following is a consistent specification of the parameters? |
| A | $\lambda = 2, a = 4, r = 3, k = 3$ |
| В | $\lambda = 4, a = 2, r = 3, k = 3$ |
| С | $\lambda = 3$, $a = 4$, $r = 2$, $k = 2$ |
| D | $\lambda = 2, a = 3, r = 2, k = 3$ |
| Answer | A |

| Id | 27 |
|----------|--|
| Question | A 2 ⁿ design in which some effects have been totally confounded is: |
| A | Always a connected design |
| В | Sometimes a connected and sometimes a disconnected design depending on the effects |
| | confounded |
| С | Always a disconnected design |
| D | Always the union of a connected and disconnected design |
| Answer | В |

| Id | 28 |
|----------|--|
| Question | A simple random sample of 25 unit is being drawn from a population of 500 units, draws |
| | being without replacement. The probability that the first population unit will be drawn in |
| | the 25 th draw is: |
| A | Zero |
| В | 1/500 |
| С | 25/500 |
| D | 475/500 |
| Answer | В |

| Id | 29 |
|----------|--|
| Question | For a given a finite population and specified sample size n. Let D1, D2, D3 respectively |
| | denote SRSWR, SRSWOR and Stratified random sampling (with Proportional allocation |
| | schemes). These can be placed in the ascending order of sampling variability: |
| A | D_1, D_2, D_3 |
| В | D_3, D_1, D_2 |
| С | D_3, D_2, D_1 |
| D | D_2, D_3, D_1 |
| Answer | С |

| Id | 30 |
|----------|---|
| Question | A finite population of N-5 units has mean $\overline{Y_N} = 12$, $S^2 = 100$. A simple random |
| | sample of n=2 units is drawn without replacement and the sample mean is denoted by |
| | $\overline{y_n}$, then $E(\overline{y_n^2})$ is |
| A | 30 |
| В | 50 |
| С | 144 |
| D | 174 |
| Answer | D |

| Id | 31 |
|----------|---|
| Question | The total number of samples of size 2 that can be drawn with replacement from a |
| | population of 10 units is |
| A | 10^2 |
| В | 2^{10} |
| С | 90 |
| D | 45 |
| Answer | A |

| Id | 32 |
|----------|--|
| Question | The regression estimator is equally efficient to the mean per unit estimator if: |
| A | $0 < \rho < 1$ |
| В | $\rho = 0$ |
| С | $-1 < \rho < 0$ |
| D | $\rho = 1$ |
| Answer | В |

| Id | 33 |
|----------|---|
| Question | Let x_1, x_2, \dots, x_n be a random sample from $U(0, \theta)$. Let $\delta(x) = 2\bar{x}$. Which of |
| | the following statement is true? |
| A | $\delta(x)$ is biased & $V(\delta(x)) > \frac{\theta^2}{n}$ |
| В | $\delta(x)$ is biased & $V(\delta(x)) < \frac{\theta^2}{n}$ |
| С | $\delta(x)$ is unbiased & $V(\delta(x)) > \frac{\theta^2}{n}$ |
| D | $\delta(x)$ is unbiased & $V(\delta(x)) < \frac{\theta^2}{n}$ |
| Answer | D |

| Id | 34 |
|----------|--|
| Question | A random sample of size 4 is drawn from normal distribution with unknown mean µ |
| | and variance σ^2 known but unspecified. For testing $H_0: \mu=0$ v/s $H_1: \mu>0$. The |
| | best critical region of the test is given by $X_1+X_2+X_3+X_4 \ge 19.6$. Which of the |
| | following values of σ should be specified so that the significance level (size) of this |
| | test is 0.025? |
| A | 5 |
| В | 15 |
| C | 10 |
| D | 20 |
| Answer | A |

| Id | 35 |
|----------|---|
| Question | Suppose X is random variable with p.d.f. |
| | $f(x;\theta) = \begin{cases} (\theta+1)x^{\theta}; 0 < x < 1\\ 0; otherwise \end{cases}$ |
| | The hypothesis $H_0: \theta=1$ is rejected in favour of $H_1: \theta=2$ if $x>0.90$. What is the |
| | probability of Type-I error? |
| A | 0.05 |
| В | 0.095 |
| С | 0.81 |
| D | 0.19 |
| Answer | D |

| Id | 36 |
|----------|---|
| Question | If for a given α , $0 \le \alpha \le 1$, non-randomized NP & LRT of a simple hypothesis V/s |
| | simple alternative exists then |
| A | They are equivalent |
| В | They are one and the same |
| С | They are exactly opposite |
| D | One can't say anything about it |
| Answer | A |

| Id | 37 |
|----------|--|
| Question | Mann- Whitney test statistic U depends on the fact that: |
| A | How many times Yi's precede Xi's |
| В | How many times Xi's precede Yi's' |
| С | Both (A) and (B) |
| D | None of (A) and (B) |
| Answer | C |

| Id | 38 |
|----------|--|
| Question | The statistic H under the Kruskal-Wallis test is approximately distributed as: |
| A | Student t |
| В | Snedecor's F |
| С | Chi-square |
| D | Normal deviate-Z |
| Answer | C |

| Id | 39 |
|----------|--|
| Question | Randomness of a sequence through runs test is adjusted by comparing the observed |
| | number of runs with: |
| A | Two critical values |
| В | One upper critical value |
| С | One lower critical value |
| D | None of the above |
| Answer | A |

| Id | 40 | | |
|----------|--|------------------------|-----|
| Question | Let the regression lines of Y on X and X on Y are respectively | $Y = \alpha X + \beta$ | and |
| | $X = \theta Y + \delta$. Then the ratio of the variances of X and Y is: | | |
| A | $\frac{\theta}{\alpha}$ | | |
| В | $\sqrt{\frac{\theta}{\alpha}}$ | | |
| С | $\sqrt{\theta \alpha}$ | | |
| D | $\frac{\alpha}{\theta}$ | | |
| Answer | A | | |

| Id | 41 |
|----------|--|
| Question | If an unknown potential parameter of a model is equalised to one or more parameters of |
| | the model, then the parameters of the model, then the parameter is classified as: |
| A | Nuisance parameter |
| В | Constrained parameter |
| С | Free parameter |
| D | Non-Centrality parameter |
| Answer | В |

| Id | 42 |
|----------|---|
| Question | The Shewhart control charts are meant: |
| A | To detect whether the process is under statistical quality control. |
| В | To find the assignable causes |
| C | To reflect the selection of samples. |
| D | All the above |
| Answer | D |

| Id | 43 |
|----------|--|
| Question | Hotellings T^2 - test is a generalization of |
| A | Chi-square test |
| В | F- test |
| С | T- test |
| D | Likelihood ratio test |
| Answer | C |

| Id | 44 |
|----------|---|
| Question | Let \underline{X} be a p-component random vector with $E[\underline{X}] = \underline{0}$ and variance-covariance |
| | matrix Σ , positive definite. If \underline{X} is partitioned into $\underline{X}^{(1)}$ of p_1 components |
| | and $X^{(2)}$ of p_2 components such that $p_1 + p_2 = p$ and $p_1 \le p_2$. Then, the square |
| | of canonical correlation are the characteristic roots of the matrix: |
| A | Σ |
| В | $\sum_{11}^{-1} \sum_{12}$ |
| С | $\sum_{22}^{-1} \sum_{21}$ |
| D | $\sum_{11}^{-1} \sum_{12} \sum_{22}^{-1} \sum_{21}$ |
| Answer | D |

| Id | 45 |
|----------|--|
| Question | The simplest linear regression modal is $Y = \alpha + \beta X + \epsilon$ where $\epsilon \sim N(0, \sigma^2)$. Let x^* |
| | denote a particular value of the independent variable X. Which of the following statement |
| | is true regarding positive variance of Y where X=x*? |
| A | $\sigma_{y x^*}^2 = \sigma^2$ |
| В | $\sigma_{y x^*}^2 = \alpha + \beta x^* + \sigma_{\epsilon}^2$ |
| С | $\sigma_{y x}^2 = \beta x^* + \sigma^2$ |
| D | $\sigma_{y x^*}^2 = \beta^2 \sigma^2$ |
| Answer | A |

| Id | 46 |
|----------|--|
| Question | Consider the simple linear regression model: $Y_i = \theta_0 + \theta_1 x_i + \epsilon_i$, $i = 1,2$ where |
| | $\epsilon_i \sim NZD(0, \sigma^2)$, $x_1 = -1$, $x_2 = 1$. The BLUE of θ_0 & θ_1 respectively are |
| A | $\left[\bar{Y}, \frac{Y_2 - Y_1}{2}\right]$ |
| В | $\left[\frac{Y_2 - Y_1}{2}, \bar{Y}\right]$ |
| С | $[Y_1 - \overline{Y}, Y_2 - \overline{Y}]$ |
| D | $\left[ar{Y},ar{Y} ight]$ |
| Answer | В |

| Id | 47 |
|----------|--|
| Question | Let $X_t = Z_t + \theta Z_{t-1}$ and $Y_t = Z_t + \frac{1}{\theta} Z_{t-1}$ where $Z_t \sim iid\ Normal(0,1)$. Which of |
| | the following statements is true? |
| A | ACF of $\{X_t\}$ process is same as that of $\{Y_t\}$ |
| В | ACF of $\{X_t\}$ process is different as that of $\{Y_t\}$ |
| С | Both the processes $\{X_t\}$ and $\{\{Y_t\}\}$ are invertible for a given value of θ |
| D | Both $\{X_t\}$ and $\{Y_t\}$ are non-stationary |
| Answer | A |

| Id | 48 |
|----------|-------------------------------------|
| Question | Polynomial trend generally follows: |
| A | T-shaped |
| В | U-shape |
| С | Linear |
| D | V-shape |
| Answer | В |

| Id | 49 |
|----------|--|
| Question | In the following regression equation: $Y_t = \alpha + \beta t + \epsilon$ where, t-time index, Y_t - |
| | number of motorcycle riders at time t, and t= 1,2,3, which of the following component |
| | from above regression equation leads to trend? |
| A | t |
| В | ε |
| С | α |
| D | β |
| Answer | D |

| Id | 50 |
|----------|---|
| Question | Consider a stationary time series Y_t given by $Y_t - \phi Y_{t-1} = \mu + Z_t - \theta Z_{t-1}$ |
| | Where Z_t is a unit noise process $(0, \sigma^2)$. Which of the following statement(s) is |
| | (are) true? |
| | 1. $E(Y_t) = \mu, V(Y_t) = \frac{\sigma^2}{(1 - \phi^2)}$ |
| | 2. $E(Y_t) = \frac{\mu}{1-\theta}, V(Y_t) = \frac{\sigma^2}{(1-\phi^2)}$ |
| | 3. $E(Y_t) = \frac{\mu}{1-\theta}$, $V(Y_t) = \frac{1-2\phi\theta+\theta^2}{(1-\theta^2)}\sigma^2$ |
| | 4. $E(Y_t) = \frac{\mu}{1-\phi}$, $V(Y_t) = \frac{1-2\phi\theta+\theta^2}{(1-\theta^2)}\sigma^2$ |
| | Codes: |
| A | (1), (2), (3) and (4) |
| В | (2) & (4) only |
| С | (4) only |
| D | (1), (2), (3) |
| Answer | C |